

A comparison between two healthy diet scores, the modified Mediterranean diet score and the Healthy Nordic Food Index, in relation to all-cause and cause-specific mortality

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Abstract

High adherence to healthy diets has the potential to prevent disease and prolong life span, and healthy dietary pattern scores have each been associated with disease and mortality. We studied two commonly promoted healthy diet scores (modified Mediterranean diet score (mMED) and the Healthy Nordic Food Index (HNFI)) and the combined effect of the two scores in association with all-cause and cause-specific mortality (cancer, CVD and ischaemic heart disease). The study included 38 428 women (median age of 61 years) from the Swedish Mammography Cohort. Diet and covariate data were collected in a questionnaire. mMED and HNFI were generated and categorised into low-, medium- and high-adherence groups, and in nine combinations of these. Multivariable-adjusted hazard ratios (HR) of register-ascertained mortality and 95% CI were calculated in Cox proportional hazards regression analysis. During follow-up (median: 17 years), 10 478 women died. In the high-adherence categories compared with low-adherence categories, the HR for all-cause mortality was 0.76 (95% CI 0.70, 0.81) for mMED and 0.89 (95% CI 0.83, 0.96) for HNFI. Higher adherence to mMED was associated with lower mortality in each stratum of HNFI in the combined analysis. In general, mMED, compared with HNFI, was more strongly associated with a lower cause-specific mortality. In Swedish women, both mMED and HNFI were inversely associated with all-cause and cardiovascular mortality. The combined analysis, however, indicated an advantage to be adherent to the mMED. The present version of HNFI did not associate with mortality independent of mMED score.

Keywords: Modified Mediterranean diet score: Healthy Nordic Food Index: Mortality: CVD: Cohort studies

Dietary risks, including high Na intake and low consumption of fruits, vegetables and whole grains, have been identified among the most important risk factors for the global burden of disease⁽¹⁾. Thus, higher adherence to a healthy diet has the potential to prevent disease and prolong life span and is emphasised in current dietary guidelines^(2,3). Healthy dietary patterns comprise different food combinations with common core foods influenced by regional food culture^(4,5). Diet scores or empirically derived food patterns are used to study healthy dietary patterns in a population, both with strengths and weaknesses⁽⁶⁾. Diet scores are often influenced by official dietary recommendations – for example, the Diet Quality Index and the Healthy Eating Index⁽⁷⁾ – but also the regional context⁽⁸⁾. The Mediterranean diet is the most investigated healthy dietary pattern, and a high adherence to such a diet has been linked to less chronic metabolic disease, longer survival and favourable cardiometabolic risk markers in observational studies^(9,10). The Lyon diet heart (secondary prevention)⁽¹¹⁾ and the Prevencion con Dieta Mediterranea (PREDIMED) (primary prevention)⁽¹²⁾ trials have also showed beneficial effects of a

Mediterranean diet on cardiovascular end points. A recent systematic review and meta-analysis on clinical trials supported favourable effects of the Mediterranean diet on cardiovascular outcomes, although the investigators raised a concern that the ‘quantity and quality of the available evidence is relatively limited’⁽¹³⁾. Many observational studies investigating Mediterranean diet scores and disease risk have been conducted in non-Mediterranean countries with other food cultures, which influence the level of intakes and thus scoring of components and classification of adherent participants. This may influence association with disease and comparability between studies⁽⁸⁾. However, this does not necessarily disqualify the generation of a Mediterranean type of diet score in a non-Mediterranean context⁽¹⁴⁾, but is one way of assessing a prudent diet in line with contemporary dietary recommendations^(2,3). Alternately, a regional-specific dietary pattern that is more reflective of foods from the Nordic countries might be better to use in a Swedish context. The Nordic diet has been linked to amelioration of cardiovascular risk factors such as inflammatory markers and lipids in intervention studies^(15,16). Different scores based on

Abbreviations: HNFI, Healthy Nordic Food Index; mMED, modified Mediterranean diet score.

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Nordic foods have been developed: the Healthy Nordic Food Index (HNFI) that was developed in Denmark⁽¹⁷⁾ and adapted for a Swedish population⁽¹⁸⁾, the Baltic Sea Diet Score (BSDS) from Finland⁽¹⁹⁾ and the New Nordic Diet score (NNDS) from Norway⁽²⁰⁾. The only score that is purely based on food items is the HNFI. BSDS includes information on fat intake, whereas the NNDS takes meal patterns into account. In cohort studies, high adherence to different Nordic diet scores has been associated with a general healthy diet composition^(19,21) and positive health outcomes^(17,18,22). However, not all reported associations have been favourable; the HNFI was not associated with CVD in one analysis⁽²³⁾ and the association between BSDS and cardiometabolic risks was deemed unclear⁽²⁴⁾.

In the present cohort study in women, we aim to study the association of two healthy diet scores: a modified Mediterranean diet score (mMED) adapted to be suitable in a non-Mediterranean context and the HNFI, as well as the joint effect of the two scores with all-cause and cause-specific mortality in a Swedish setting. A high adherence to each score is reflective of adherence to a healthy diet, although defined in different ways. Because of the large number of outcomes, it will, for the first time, be possible to investigate mortality in women highly adherent to one diet but not to the other.

Methods

Study population

The Swedish Mammography Cohort was established in 1987–1990. All 90 303 women residing in two Swedish counties, born between

1914 and 1948, were invited to a mammography screening. Enclosed with this invitation was a questionnaire covering diet and lifestyle, which was completed by 74% of the women. In late 1997, a second expanded questionnaire was sent to all participating women still living (*n* 56 030; 70% of those eligible) in the study. The study sample with exclusions has been described previously⁽²⁵⁾. For the present analysis, the 1997 investigation was considered as baseline. To define the study sample (Fig. 1), we excluded individuals with implausible energy intakes ± 3 SD from the ln-transformed mean total energy intake (*n* 521) and thirty-five individuals who completed the questionnaire but died before 1 January 1998, leaving 38 428 for analysis. Those individuals with missing values of one or more components of the diet score were kept in the data set and treated as null reporting, but were excluded in sensitivity analysis. Missing values were <5% for all food groups used to define the scores, except for oatmeal porridge, which had 14% missing values. In the analysis of cause-specific mortality, participants who had been diagnosed with cancer and CVD (*n* 5087) before the investigation in 1997 were excluded, leaving 33 341 for analysis. The study was approved by the regional ethics committees at Uppsala University, Uppsala, and Karolinska Institutet, Stockholm, Sweden.

Dietary and covariate assessments

The FFQ has been described in more detail previously⁽²⁶⁾. In short, the FFQ included ninety-six food items and the participants indicated in the FFQ how often, on average, they had consumed different food items during the past year and chose from eight pre-defined frequency categories ranging from 'never/seldom'

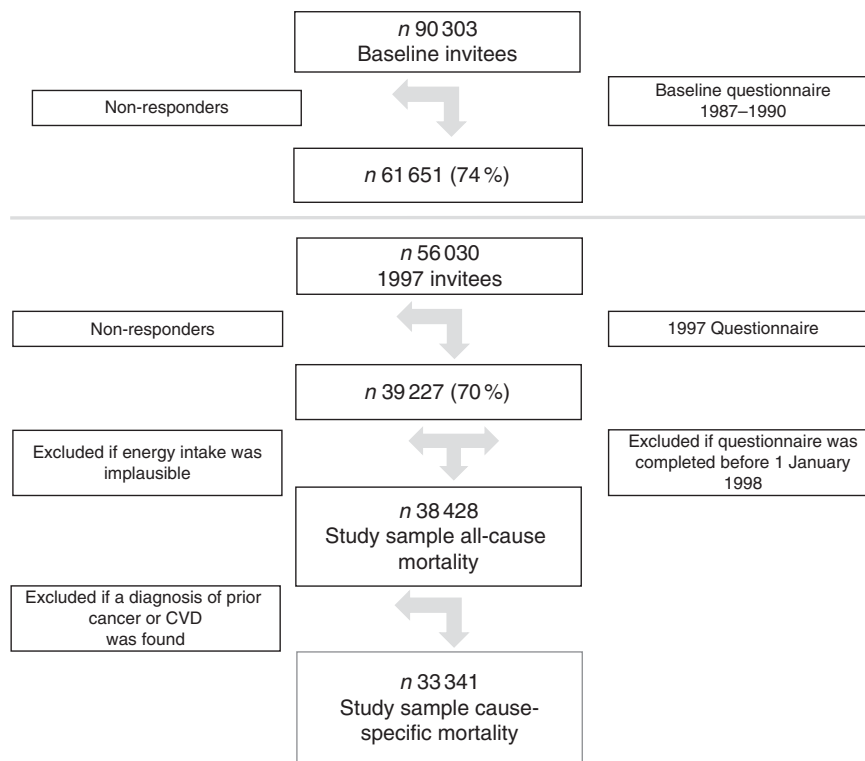


Fig. 1. The flow chart depicts the study sample with exclusions. The baseline for the present analyses was the 1997 investigation and the study sample used in the analysis with all-cause mortality was 38 428 and 33 341 for cause-specific mortality.

to '3 or more times per day'. Frequently consumed foods such as dairy products and bread slices were additionally reported as number of servings per day. Information on fat type used in cooking and as salad dressing was also reported. Total amount of alcohol consumed per day was derived from the FFQ by multiplying the reported frequency with the reported amount on a single occasion. Energy and nutrient intakes were estimated by multiplying the consumption frequency of each food item with the nutrient content (including energy) of age-specific portion sizes. Nutrient values were obtained from the Swedish food composition database, National Food Agency. Nutrient intakes were adjusted for total energy intake using the residual method⁽²⁷⁾. The covariates collected in the 1997 questionnaire included educational level, living alone, physical activity, smoking habits, weight (kg) and height (m). Physical activity had five pre-defined levels ranging from 1 h to more than 5 h/week, and this question has been found to be valid compared with activity records and accelerometer data⁽²⁸⁾. Energy intake and Charlson's comorbidity index were also included as covariates. Charlson's weighted comorbidity index^(29,30) was defined by ICD diagnosis codes (versions 8, 9 and 10) from in-patient care registered in the National Patient Register. BMI was calculated as weight divided by height squared.

Dietary exposures

Modified Mediterranean diet score. Adapted from the Mediterranean diet scale by Trichopoulou *et al.*⁽³¹⁾, a mMED (range 0–8 points) was calculated by use of previously defined food items⁽²⁶⁾. One point was given for intakes above the median of each of the following components: fruit and vegetables (apple, banana, berry, orange/citrus and other fruit; carrot, beetroot, broccoli, cabbage, cauliflower, lettuce, onion, garlic, pepper, spinach, tomato and other vegetables), legumes (peas, lentils, beans and pea soup) and nuts, non-refined or high-fibre grains (whole meal bread, crisp bread, oatmeal and bran of wheat), fermented dairy products (sour milk, yogurt and cheese) and fish (excluding shellfish). In addition, one point was

given for intakes below the median of red and processed meat. Any use of olive or rapeseed oil for cooking or as dressing and moderate alcohol consumption (intake range 5–15 g ethanol/d) also rendered 1 point. Median intakes and distribution of the components are presented in Table 1. A three-level categorical diet score was formed, with 0–2 points indicating low adherence, 3–5 points indicating medium adherence and 6–8 points indicating high adherence, as previously suggested⁽³²⁾.

Healthy Nordic Food Index

The HNFI (range 0–6 points) was calculated as previously done in a different Swedish population⁽¹⁸⁾, but using the frequency per day of the different foods rather than grams per day to be comparable with the mMED. The foods in the HNFI were originally chosen because of expected beneficial health effects and that the foods were originating from Nordic nature, as well as being quantitative important foods in the Nordic diet⁽¹⁷⁾. The median intakes and distribution of the six components are presented in Table 1. One point was given for intakes above the median of each of the following foods: apples and pears, root vegetables (carrot and beet root), cruciferous vegetables (broccoli, cabbage and cauliflower), whole-grain bread (whole meal bread, crisp bread of rye), oatmeal porridge and fish (salmon, mackerel, herring and white fish such as cod and shellfish). A three-level categorical diet score was formed from the HNFI, with 0–1 points indicating low adherence, 2–4 points indicating medium adherence and 5–6 points indicating high adherence, making the low and the high categories approximately similar in size. As a sensitivity analysis, the HNFI was also generated with grams per day. The generated three-level categorical variables ranked participants exactly the same using either frequencies or grams per day.

Outcomes

We considered mortality collected from the Swedish cause of death registry between study baseline 1 January 1998 and

Table 1. Distribution and percentages for foods, food groups, the Healthy Nordic Food Index and the modified Mediterranean diet score in the Swedish Mammography Cohort (Mean values and 25th (p25), 50th (p50) and 75th (p75) percentiles; *n* 38 428)

Frequency per day	Mean	p25	p50*	p75
Fruit and vegetables	4.8	3.0	4.4	6.1
Legumes and nuts	0.29	0.13	0.20	0.35
Fish	0.31	0.13	0.21	0.35
Red and processed meat	1.1	0.63	0.98	1.4
Non-refined or high-fibre grains	3.3	2.0	3.0	4.1
Fermented dairy products	3.6	2.0	3.0	4.7
Total mMED (points)	3.9	3	4	5
Whole meal and crisp bread	3.0	2.0	2.9	4.0
Fish, including shellfish	0.52	0.27	0.41	0.63
Root vegetables	0.53	0.28	0.43	0.71
Cruciferous vegetables	0.44	0.20	0.35	0.63
Apples and pears	0.64	0.21	0.5	1.0
Oatmeal porridge	0.27	0	0.07	0.5
Total HNFI (points)	2.8	2	3	4
Alcohol intake between 5 and 15 g/d		27.3%		
Any use of olive or rapeseed oil†		46.6%		

* The median intake as presented above was the cut-off used for the scoring but without rounding.

† Used in cooking and for dressings.

31 December 2014. A complete linkage with the register is possible as all Swedish residents have a personal identity number. Since 1952, the National Board of Health and Welfare has collected information on the causes of death for all Swedish residents in the cause of death registry. We used the underlying cause of death to define mortality from all causes, CVD (International Classification of Diseases, 10th revision (ICD-10) codes I00–I99), ischaemic heart disease (ICD-10 codes I20–I25) and cancer (ICD-10 C-codes).

Statistical analysis

For each participant, follow-up time was accrued from 1 January 1998 until date of death or the end of the study period (31 December 2014). The associations between mMED and HNFI on the one hand and all-cause mortality and cause-specific mortality on the other were estimated as age- and multivariable-adjusted hazard ratios (HR) and 95% CI by Cox proportional hazards regression using age as the primary time scale. Both of the diet scores – mMED and HNFI – were initially treated as continuous variables, to assess how every one-point increment in each score was related to mortality. The scores were then treated as categorical variables to assess how high, moderate and low adherence to respective dietary score was associated with mortality. The low and high categories roughly reflected the lowest and the highest quintiles. We also investigated the trend over categories. The combination of mMED and HNFI was thereafter used to jointly classify study participants into nine categories. The high mMED/high HNFI level was the reference category in the analyses. The combined analysis was performed for all-cause mortality but not for cause-specific death because of limited number of cases in some of the cells. Covariates for the present analyses were chosen using directed acyclic graphs⁽³³⁾, and the multivariable model I included educational level (≤ 9 , 10–12, >12 years, other), living alone (yes or no), physical activity (five categories), energy intake (continuous), smoking habits (current, former, never) and Charlson's weighted comorbidity index (continuous; 1–16). We further included the other diet score (HNFI or mMED) in an additional model (model II) and stratified the analysis of one diet score on every category of the other. Missing data on covariates were imputed using Stata's 'mi' package (chained multiple imputation) and twenty imputations were done to reduce sampling error. The multiple imputation takes into account model variables and produces twenty separate data sets. The Cox analysis is subsequently run on all the separate data sets and the results are combined. We imputed missing observations for physical activity level, living alone and smoking status (missing values <14%). The proportional hazard assumption in the Cox models was confirmed graphically by log-log plots of survival. As a sensitivity analysis, the main analysis was restricted to individuals with complete case data. We further re-ran the analysis excluding participants with missing data on any of the components of respective diet score. BMI was not primarily considered as a confounder but rather an intermediate, but was added to model I as a sensitivity analysis. Non-fermented milk intake has previously been positively associated with mortality in the present cohort⁽³⁴⁾ and was also added to model I as a sensitivity analysis. All analyses were carried out in Stata version 12.0 (StataCorp LP).

Results

Baseline characteristics

The baseline characteristics among the study participants stratified by adherence to mMED and HNFI are presented in Table 2. Despite modest absolute differences, mean BMI and age were the lowest in the high-adherence group of mMED, whereas an opposite pattern was observed for HNFI. Energy intake was the highest in both high-adherence groups, and there were more individuals classified as highly adherent with the mMED – 6965 individuals compared with 5527 individuals for HNFI. Nutrient intakes increased with higher adherence to either diet scores. The intake of red and processed meat was lowest in the high mMED adherence group and highest in the high HNFI adherence group. The alcohol intake levels between adherence groups of the two diet scores were opposite, with the highest mean intake in the highest mMED adherence group, whereas the lowest mean intake in the high HNFI adherence group. Further, educational level did not differ much between the different adherence categories of HNFI, whereas there was a gradient across the three adherence categories of mMED with the highest proportion of high educational attainment in the high-adherence category. For smoking status and physical activity level there was a similar tendency across the adherence categories of both mMED and HNFI.

Diet scores and mortality

During a median of 17 years of follow-up and 583 826 person-years at risk, 10 478 women died among the 38 428 women who entered the study. After exclusion of individuals with prior cancer or CVD, 33 341 women remained to follow-up for cause-specific death; of those, 2355 died from cancer, 3003 from cardiovascular causes and 1081 from ischaemic heart disease, with a total of 518 031 person-years at risk.

Age- and multivariable-adjusted HR for all-cause and cause-specific mortality are presented in Tables 3 and 4, respectively. For all-cause mortality, those classified in the highest compared with the lowest category of the two diet scores had lower mortality rates. The multivariable-adjusted HR (model I) for all-cause mortality with a high adherence of mMED, compared with low adherence, was 0.76 (95% CI 0.70, 0.81). Comparing high adherence of HNFI with low adherence, the HR for all-cause mortality was 0.89 (95% CI 0.83, 0.96). Although we detected an educational gradient across mMED categories, we obtained similar estimates stratified on educational attainment: 0.76 (95% CI 0.70, 0.83) among those with ≤ 9 years of schooling and 0.70 (95% CI 0.56, 0.86) among those with >9 years of schooling in *post hoc* analysis. Further, to examine whether the detected differences in certain food intakes over adherence categories of HNFI influenced the estimates, we added the intake of red and processed meat and alcohol to model I, and this did not attenuate the estimates.

After mutual adjustment for the other diet score (model II), the inverse association remained for mMED (0.76; 95% CI 0.82, 0.90), comparing the highest with the lowest adherence categories, whereas it was attenuated for HNFI. This was also confirmed when the analysis with respective diet score was stratified on every category of the other diet score (Table 5). The strongest inverse association between mMED and all-cause mortality was

Table 2. Characteristics of study participants according to adherence category of the modified Mediterranean diet score (mMED) and the Healthy Nordic Food Index (HNFI) taking part of the Swedish Mammography Cohort (Numbers and percentages; mean values and standard deviations)

Adherences	mMED						HNFI					
	Low (0–2 points)		Medium (3–5 points)		High (6–8 points)		Low (0–1 points)		Medium (2–4 points)		High (5–6 points)	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
Number of subjects	7992		23 469		6965		8195		24 704		5527	
Age at study entry (years)	63	9.6	62	9.3	61	8.7	61	9.2	62	9.3	63	9.1
BMI (kg/m ²)	25.4	4.2	25.1	4.0	24.5	3.7	24.9	4.0	25.0	3.9	25.3	4.0
Educational level												
Other												
<i>n</i>	30		60		14		25		65		14	
%	0.4		0.3		0.2		0.3		0.3		0.3	
< 9 years												
<i>n</i>	6780		17 449		4 203		6093		18 150		4189	
%	84		74		60		74		73		76	
Between 9 and 12 years												
<i>n</i>	441		1771		653		645		1838		382	
%	5.5		7.5		9.4		7.9		7.5		6.9	
>12 years												
<i>n</i>	742		4190		2095		1434		4651		942	
%	9.2		18		30		18		19		17	
Smoking status												
Current												
<i>n</i>	2164		5263		1287		2542		5270		902	
%	27		22		18		31		21		16	
Former												
<i>n</i>	1508		5223		1929		1861		5576		1223	
%	19		22		28		23		23		22	
No												
<i>n</i>	4149		12 560		3651		3625		13 420		3315	
%	52		54		52		44		54		60	
Missing												
<i>n</i>	172		424		98		169		438		87	
%	2.2		1.8		1.4		2.1		1.8		1.6	
Physical activity level												
1 (Low)												
<i>n</i>	1878		4125		809		2084		4149		579	
%	23		18		12		25		17		10	
2												
<i>n</i>	1621		5018		1463		1781		5272		1049	
%	20		21		21		22		21		19	
3												
<i>n</i>	2052		7097		2367		2100		7549		1867	
%	26		30		34		26		31		34	
4												
<i>n</i>	648		2488		925		600		2686		775	
%	8.1		11		13		7.3		11		14	
5 (High)												
<i>n</i>	620		2319		895		569		2473		792	
%	7.8		10		13		6.9		10		14	
Missing												
<i>n</i>	1174		2423		506		1063		2575		465	
%	15		10		7.2		13		10		8.4	
Residual adjusted nutrient intakes per day												
Vitamin C (mg)	96	52	114	57	131	52	92	49	116	57	135	54
Fibre (g)	19	5.1	22	5.5	25	5.0	18	4.6	23	5.1	26	5.0
Fe (mg)	11	3.5	11	3.0	11	2.4	10	2.9	11	3.0	12	3.0
Ca (mg)	994	334	1047	298	1087	259	1074	344	1037	293	1025	259
Sucrose (g)	41	21	35	16	32	12	39	21	36	16	33	13
Vitamin D (µg)	4.4	1.7	4.5	1.7	4.6	1.6	4.2	1.4	4.5	1.7	4.9	1.7
Protein (E%)	15.9	2.8	16.7	2.7	17.3	2.6	16.3	2.8	16.6	2.7	17.3	2.7
Saturated fat (E%)	14.6	3.5	14.4	3.4	14.0	3.2	15.5	3.6	14.2	3.3	13.3	3.0
Alcohol (E%)	1.4	2.4	1.7	2.2	2.1	1.7	2.1	2.6	1.7	2.0	1.3	1.5
Energy (kJ)	5849	1799	7355	2092	8552	2092	6050	1841	7330	2050	8732	2343
Energy (kcal)	1398	430	1758	500	2044	500	1446	440	1752	490	2087	560

Table 2. *Continued*

Adherences	mMED						HNFI					
	Low (0–2 points)		Medium (3–5 points)		High (6–8 points)		Low (0–1 points)		Medium (2–4 points)		High (5–6 points)	
	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
Food intake in frequency per day												
Non-refined cereals	2.2	1.5	3.4	2.0	4.2	1.9	2.0	1.4	3.4	1.9	4.7	1.9
Red and processed meat	1.1	0.7	1.1	0.8	1.0	0.7	0.9	0.6	1.1	0.7	1.3	0.9
Fruit and vegetables	3.0	1.8	4.9	2.6	6.8	2.7	2.9	1.6	5.0	2.5	7.2	3.1
Fish and shellfish	0.4	0.4	0.5	0.5	0.7	0.5	0.3	0.3	0.5	0.5	0.8	0.6
Fermented dairy products	0.9	0.8	1.1	1.0	1.3	1.1	1.0	1.0	1.1	1.0	1.2	1.0
Legumes and nuts	0.2	0.2	0.3	0.3	0.4	0.3	0.2	0.2	0.3	0.3	0.4	0.4
Root vegetables	0.4	0.4	0.5	0.5	0.7	0.5	0.2	0.2	0.6	0.4	0.9	0.5
Apples and pears	0.5	0.5	0.6	0.6	0.8	0.6	0.3	0.4	0.7	0.6	0.9	0.6
Oatmeal porridge	0.2	0.3	0.3	0.4	0.3	0.4	0.1	0.2	0.3	0.4	0.5	0.4
Cruciferous vegetables	0.3	0.3	0.4	0.5	0.6	0.6	0.2	0.2	0.4	0.5	0.8	0.6

E%, percentage of energy intake.

Table 3. All-cause mortality stratified by adherence to respective diet (modified Mediterranean diet score (mMED) and Healthy Nordic Food Index (HNFI)) score in the Swedish Mammography Cohort (Age- and multivariable-adjusted hazard ratios (HR)* and 95 % confidence intervals; *n* 38 428)

Adherences	Low		Medium		High		Trend per adherence category		Per 1 point increase	
	HR	95 % CI	HR	95 % CI	HR	95 % CI	HR	95 % CI	HR	95 % CI
mMED										
Score points	0–2		3–5		6–8					
Number of participants	7993		23 470		6965					
Number of deaths	2706		6365		1407					
Person-years at risk	116 7276		357 1683		109 9303					
Rate per 1000 person-years	23.2	22.3, 24.0	17.8	17.4, 18.3	12.8	12.1, 13.5				
Age adjusted	1.0	Ref.	0.85	0.81, 0.88	0.71	0.66, 0.75	0.84	0.82, 0.87	0.93	0.92, 0.94
Adjusted (Model I)†	1.0	Ref.	0.86	0.82, 0.90	0.76	0.70, 0.81	0.87	0.84, 0.90	0.94	0.93, 0.95
Adjusted (Model II)‡	1.0	Ref.	0.87	0.82, 0.91	0.76	0.82, 0.90	0.87	0.82, 0.90	0.94	0.92, 0.95
HNFI										
Score points	0–1		2–4		5–6					
Number of participants	8197		24 704		5527					
Number of deaths	2179		6708		1591					
Person-years at risk	124 2128		375 5950		84 0185					
Rate per 1000 person-years	17.5	16.8, 18.3	17.9	17.4, 18.3	18.9	18.0, 19.9				
Age adjusted	1.0	Ref.	0.87	0.83, 0.91	0.83	0.78, 0.88	0.91	0.88, 0.94	0.96	0.95, 0.97
Adjusted (Model I)†	1.0	Ref.	0.91	0.87, 0.96	0.89	0.83, 0.96	0.94	0.91, 0.98	0.98	0.96, 0.99
Adjusted (Model II)‡	1.0	Ref.	0.96	0.91, 1.00	0.98	0.91, 1.06	0.98	0.95, 1.02	1.00	0.99, 1.02

Ref., referent values.

* HR were determined in Cox proportional hazard analysis.

† Model I was adjusted for educational level (≤ 9 , 10–12, >12 years, other), living alone (yes or no), physical activity (five categories), energy intake (continuous), smoking habits (current, former, never) and Charlson's comorbidity index (continuous; 1–16).

‡ Model II was adjusted for model I and the other diet score (mMED or HNFI).

observed in the low adherence category of HNFI. Fig. 2 illustrates the combined exposure of mMED and HNFI on mortality and shows that higher adherence to the mMED was associated with lower all-cause mortality in each stratum of the HNFI.

For cause-specific mortality, the associations with mMED were in general stronger than with HNFI. Cancer mortality was not independently associated with HNFI, whereas higher adherence of the mMED was associated with lower cancer mortality in a dose–response manner by adherence category (model I). For mortality due to CVD, there was an inverse association for both diet scores, but the association was abolished for HNFI when adjusting for mMED. The inverse age-adjusted association between the HNFI and mortality due to ischaemic heart disease was attenuated after adjustment for covariates, while the association remained for mMED.

Complete case analysis (all-cause mortality cohort 29 415 and 7811 deaths; cause-specific death cohort 25 754 and 1829 cancer deaths, 2276 cardiovascular deaths and 812 ischaemic heart disease deaths) confirmed the reported associations for all-cause and cause-specific mortality (data not shown). When excluding individuals with missing data on any of the score components of the diet scores, the estimates did not materially change for neither all-cause mortality nor cause-specific mortality. Adding BMI or non-fermented milk intake to model I did not change estimates either.

Discussion

In the present study, two scores reflective of a healthy diet, mMED and HNFI, were inversely associated with all-cause and

Table 4. Cause-specific death due to cancer, CVD and ischaemic heart disease stratified by adherence to modified Mediterranean diet score (mMED) and Healthy Nordic Food Index (HNFI) in the Swedish Mammography Cohort (Age- and multivariable-adjusted hazard ratios (HR)* and 95% CI; *n* 33 341)

Adherences	Low		Medium		High		Trend per adherence category	
	HR	95% CI	HR	95% CI	HR	95% CI	HR	95% CI
Cancer mortality (<i>n</i> 2355)								
mMED								
Number of participants	5884		19 113		5985			
Number of deaths	496		1467		392			
Person-years at risk	96 4251		319 6687		101 9375			
Rate per 1000 person-years	5.1	4.7, 5.6	4.6	4.4, 4.8	3.8	3.5, 4.2		
Age adjusted	1.0	Ref.	0.92	0.83, 1.02	0.82	0.72, 0.94	0.91	0.85, 0.97
Adjusted (Model I)†	1.0	Ref.	0.93	0.83, 1.03	0.84	0.72, 0.97	0.92	0.85, 0.99
Adjusted (Model II)‡	1.0	Ref.	0.92	0.82, 1.03	0.81	0.69, 0.94	0.91	0.84, 0.99
HNFI								
Number of participants	6447		20 058		4481			
Number of deaths	466		1497		392			
Person-years at risk	107 9292		334 4671		75 6350			
Rate per 1000 person-years	4.3	3.9, 4.7	4.4	4.3, 4.7	5.2	4.7, 5.7		
Age adjusted	1.0	Ref.	0.92	0.83, 1.02	0.98	0.85, 1.12	0.98	0.92, 1.05
Adjusted (Model I)†	1.0	Ref.	0.97	0.87, 1.08	1.06	0.91, 1.23	1.03	0.95, 1.11
Adjusted (Model II)‡	1.0	Ref.	1.0	0.89, 1.12	1.14	0.97, 1.32	1.06	0.98, 1.15
Cardiovascular mortality (<i>n</i> 3003)								
mMED								
Number of participants	5655		18 693		5990			
Number of deaths	725		1887		391			
Person-years at risk	96 4251		319 6687		101 9375			
Rate per 1000 person-years	7.5	7.0, 8.1	5.9	5.6, 6.2	3.8	3.5, 4.2		
Age adjusted	1.0	Ref.	0.84	0.73, 0.97	0.61	0.50, 0.75	0.82	0.78, 0.87
Adjusted (Model I)†	1.0	Ref.	0.86	0.74, 1.00	0.65	0.52, 0.82	0.84	0.79, 0.90
Adjusted (Model II)‡	1.0	Ref.	0.87	0.74, 1.02	0.65	0.52, 0.83	0.85	0.79, 0.91
HNFI								
Number of participants	6342		19 573		4424			
Number of deaths	571		1983		449			
Person-years at risk	107 9292		334 4671		75 6350			
Rate per 1000 person-years	5.3	4.9, 5.7	5.9	5.7, 6.2	5.9	5.4, 6.5		
Age adjusted	1.0	Ref.	0.90	0.82, 0.99	0.78	0.69, 0.88	0.88	0.83, 0.94
Adjusted (Model I)†	1.0	Ref.	0.92	0.84, 1.03	0.82	0.72, 0.94	0.91	0.85, 0.97
Adjusted (Model II)‡	1.0	Ref.	0.97	0.88, 1.07	0.91	0.79, 1.05	0.96	0.89, 1.03
Mortality due to ischaemic heart disease (<i>n</i> 1081)								
mMED								
Number of participants	6115		19 899		6246			
Number of deaths	265		681		135			
Person-years at risk	96 4251		319 6687		101 9375			
Rate per 1000 person-years	2.7	2.4, 3.1	2.1	2.0, 2.3	1.3	1.1, 1.6		
Age adjusted	1.0	Ref.	0.88	0.81, 0.96	0.77	0.69, 0.86	0.79	0.72, 0.87
Adjusted (Model I)†	1.0	Ref.	0.92	0.84, 1.00	0.83	0.73, 0.94	0.82	0.73, 0.91
Adjusted (Model II)‡	1.0	Ref.	0.97	0.88, 1.07	0.95	0.83, 1.09	0.82	0.73, 0.92
HNFI								
Number of participants	6698		20 857		4705			
Number of deaths	215		698		168			
Person-years at risk	107 9292		334 4671		75 6350			
Rate per 1000 person-years	2.0	1.7, 2.3	2.1	1.9, 2.2	2.2	1.9, 2.6		
Age adjusted	1.0	Ref.	0.85	0.73, 0.99	0.78	0.64, 0.95	0.88	0.79, 0.97
Adjusted (Model I)†	1.0	Ref.	0.89	0.76, 1.05	0.86	0.69, 1.07	0.92	0.83, 1.03
Adjusted (Model II)‡	1.0	Ref.	0.94	0.80, 1.11	0.97	0.77, 1.23	0.98	0.87, 1.11

Ref., referent values.

* HR were determined in Cox proportional hazard analysis.

† Model I was adjusted for educational level (≤ 9 , 10–12, > 12 years, other), living alone (yes or no), physical activity (five categories), energy intake (continuous), smoking habits (current, former, never) and Charlson's comorbidity index (continuous; 1–16).

‡ Model II was adjusted for model I and the other diet score (mMED or HNFI).

cardiovascular mortality, although the strength of the associations differed between the scores. Those women who were classified in the highest compared with the lowest category of the mMED had a 24% lower multivariable-adjusted all-cause mortality, whereas the corresponding estimate for HNFI was 11%.

Thus, a high scoring (high adherence) of the mMED compared with a high scoring of the HNFI in this cohort of Swedish women seems to be associated with an advantage in terms of survival. This was further confirmed in the analysis with combined exposures; a higher adherence to the mMED was associated with

Table 5. All-cause mortality in association with adherence to modified Mediterranean diet score (mMED) and Healthy Nordic Food Index (HNFI) stratified on each adherence category of the other diet score in the Swedish Mammography Cohort (Numbers and percentages; age- and multivariable-adjusted hazard ratios (HR)* and 95% CI; *n* 38 428)

	<i>n</i>	%	Low		Medium		High	
			HR	95% CI	HR	95% CI	HR	95% CI
mMED stratified on each adherence category of HNFI								
HNFI category (low)	8197	21	1.0	Ref.	0.80	0.74, 0.87	0.63	0.44, 0.90
HNFI category (medium)	24 704	64	1.0	Ref.	0.89	0.84, 0.94	0.75	0.69, 0.82
HNFI category (high)	5527	14	1.0	Ref.	0.85	0.62, 1.16	0.70	0.51, 0.97
HNFI stratified on each adherence category of mMED								
mMED category (low)	7993	21	1.0	Ref.	0.89	0.83, 0.97	0.94	0.69, 1.29
mMED category (medium)	23 470	61	1.0	Ref.	0.97	0.90, 1.04	0.99	0.90, 1.09
mMED category (high)	6965	18	1.0	Ref.	1.00	0.70, 1.44	1.00	0.69, 1.43

Ref, referent values.

* HR were determined in Cox proportional hazard analysis.

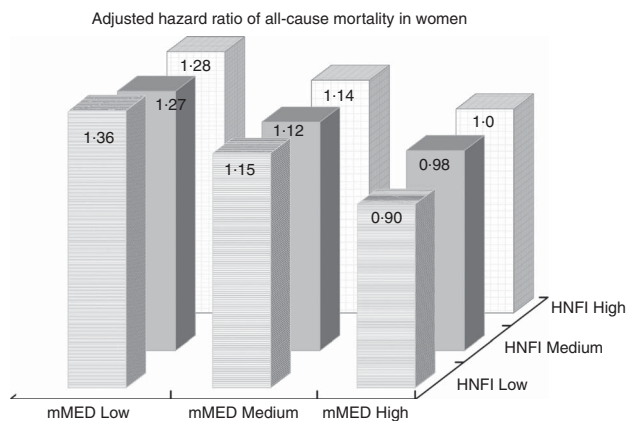


Fig. 2. Multivariable-adjusted hazard ratio of mortality in women across nine strata formed by the combined categories of modified Mediterranean diet score (mMED) and Healthy Nordic Food Index (HNFI). The high mMED/high HNFI adherence category was used as the reference category. The model was adjusted for educational level (≤ 9 , 10–12, > 12 years, other), living alone (yes or no), physical activity (five categories), energy intake (continuous), smoking habits (current, former, never) and Charlson's weighted comorbidity index (continuous; 1–16).

lower mortality in each stratum of the HNFI, whereas results for HNFI were not independent from mMED. It was previously noted that the HNFI may lack important features of a healthy Nordic diet⁽³⁵⁾, including the use of rapeseed oil, and that a Nordic diet could be more widely defined and include more items^(19,36) than those included in the proposed HNFI⁽¹⁷⁾. These factors may indeed have influenced the findings in the present study.

To our knowledge, this is the first study to classify participants in joint exposure strata reflecting the combined adherence to the mMED and the HNFI. This analysis makes it possible to discern whether there may be an advantage to be adherent to either the HNFI or the mMED in the present cohort. As the HNFI was developed to be specific to the Nordic region, it could have been expected that high adherence to the HNFI, compared with a high adherence to mMED, would be an advantage in this Swedish population. However, this was not the case – an observation that was strengthened by the analysis across the nine joint strata as presented in the Fig. 2. Moreover, in the multivariable analysis including the other diet score, the association with mortality remained for mMED but the

association between HNFI and mortality was attenuated comparing the high- and the low-adherence categories. The scores are correlated with each other (r 0.53) but they do not reflect the same dietary components. The combination of components of the HNFI did not associate with mortality independent of the mMED score. However, this does not necessarily mean that a healthy Nordic diet *per se* is less healthy than a Mediterranean type of diet; rather, it means that the index may not capture the full potential of a healthy Nordic diet and may need refinement.

The HNFI was originally developed in Denmark, a country with some similar but also some different features of food culture than Sweden^(37,38). The HNFI comprised six different foods and food groups – apples and pears, root vegetables, cruciferous vegetables, whole-grain bread, oatmeal porridge and fish – which are all healthy foods that will provide fibre, whole grains, micronutrients such as folate, vitamin C, K and Fe, *n-3* fatty acids and other bioactive compounds, which may help prevent chronic metabolic disease⁽²⁾. In the original HNFI rye bread was included, as rye bread consumption is an important feature of the Danish diet⁽¹⁷⁾. However, rye bread is not as commonly consumed in Sweden and was exchanged with whole-grain bread when adapted for a Swedish cohort⁽¹⁸⁾. In this study, we included high-fibre and whole-grain breads, as well as crisp bread.

The mMED is a more diversified diet score and is less reliant on single foods, which may have affected the association between HNFI and mortality. The HNFI was generated by the use of fourteen food items from the FFQ, whereas the mMED was based on forty-one food items. The scoring of the mMED was also dependent on a lower intake of red and processed meat, as well as on a moderate alcohol intake. The scoring of the HNFI relied only on high intakes of a limited number of healthy foods and may thus permit individuals with a less healthy total diet in the high-adherence group. Indeed, those who were highly adherent to HNFI had the highest intake of red and processed meat, as well as the lowest intake of alcohol; these factors may have influenced the results, but adjustment for these food groups did not influence estimates. However, high adherence to the mMED was more common than high adherence of HNFI. Further, a high adherence to mMED was associated with a higher educational level, but stratifying the analysis on an educational level (shorter or longer than 9 years of schooling) did not influence the point estimates in the present study.

Previously, the HNFI has been related to lower total mortality in Denmark⁽¹⁷⁾, with a 4% lower rate for each one-point higher score in men and women, comparable to the 2% lower rate per point among women in the present study. Further, several studies based on the Swedish Women's Lifestyle and Health cohort have been published with the HNFI as the exposure. In that cohort, the HNFI was related to a healthier lifestyle⁽¹⁸⁾, whereas not to the risk of CVD⁽²³⁾, but presented a 6% lower total mortality per 1 point increment, which was confined to cancer and non-CVD causes⁽²²⁾. This is seemingly in contrast to our results, where the HNFI was independently associated with lower mortality due to CVD, but not to cancer death. The differences in results may be because of a larger number of CVD cases in the present cohort and that the women were older at follow-up and had a higher baseline BMI and thus consequently were also at a higher risk for CVD⁽³⁹⁾. In addition, median intake levels of the score components were higher in the present cohort, and our score may be reflective of a more nutrient-dense diet. Moreover, a recent study from Denmark reported an inverse association between high adherence to the HNFI and risk of myocardial infarction⁽⁴⁰⁾ over 13.6 years. The fact that we failed to see an inverse adjusted association with mortality due to ischaemic heart disease may be explained by fewer cases and that the HNFI failed to account for important features of a healthy diet, especially in a Swedish setting, important in the prevention of ischaemic heart disease, as previously suggested⁽³⁵⁾. This may relate to the content of foods rich in phytochemicals, such as carotenoids and polyphenols, which are abundant in plant foods and have antioxidant activities⁽⁹⁾ or some other component of a Nordic diet such as specific types of dairy products, berries and moderate intakes of meat⁽³⁶⁾.

The inverse association between a Mediterranean type of diet and CVD and mortality has been reported previously^(10,14,31,41,42) and includes studies from non-Mediterranean countries^(14,26,32,43,44). A recent study from the UK evaluated associations between four different Mediterranean diet scores and CVD and mortality and reported inverse associations for three of the scores, but not for the score using sex-specific medians. This score was seen as too crude and not sensitive enough to evaluate adherence to a Mediterranean diet in a non-Mediterranean context⁽⁴⁴⁾. In the present study, mMED was median-derived and was sensitive enough to classify participants into distinct adherence categories, and a higher adherence was inversely associated with all-cause and cause-specific death. This may have been influenced by the wide exposure range of the different components of the mMED, as well as our large sample size. The beneficial effects of a Mediterranean type of diet have been ascribed to high-quality fat and carbohydrates, as well as a high intake of key micronutrients and other bioactive compounds such as polyphenols with antioxidant capacity^(9,10), factors known to influence lipid levels, insulin resistance, oxidative stress and inflammatory processes, which in turn are involved in the pathophysiology of CVD^(45,46). There has been criticism regarding the use of the term Mediterranean diet in a non-Mediterranean context⁽⁴⁷⁾; however, this view is not shared by all nutrition researchers, and recent systematic reviews support beneficial associations between Mediterranean diet scores used in diverse populations and different health outcomes^(41,48).

The strengths of the present study include the long follow-up time with proportional hazards and the large number of deaths ascertained through individual linkage to the national death registry. Although we adjusted for important covariates as proxies for socio-economic status, there may be an influence of residual or unmeasured confounding. As with any self-reported exposure, reporting of diet is connected to random and systematic measurement errors, and further to difficulties in estimating portion sizes and inadequacies in food composition tables. These problems are partially compensated for by exclusion of participants with the most extreme energy intakes and by adjustment for total energy intake rendering adherence to the isoenergetic principle of the exposure categories⁽⁴⁹⁾. Furthermore, the use of only one diet assessment in a study with long follow-up time is a weakness, which may also lead to attenuated estimates, as previously discussed⁽⁵⁰⁾. The study questionnaire has the ability to rank participants and has been found to be valid and reproducible, and the large study size compensates for random misclassification. These results might not apply to other ethnicities or men.

In conclusion, a higher adherence to a healthy diet, reflected by the mMED and the HNFI, was independently and inversely associated with lower total all-cause and mortality owing to cardiovascular causes, but with different strengths. The rate of cancer death and death due to ischaemic heart disease was only inversely and independently associated with mMED. There seems to be an advantage to be adherent to the mMED, rather than HNFI, in terms of survival in this cohort of Swedish women. Further, we conclude that HNFI may not necessarily capture the full potential of a Nordic diet related to health outcomes, at least in the present cohort, and may need further development.

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E. W. L. and K. M. designed the research; E. W. L., L. B., A. W. and K. M. conducted the research; E. W. L. performed statistical analysis, drafted the paper and had primary responsibility for final content; and L. B., A. W. and K. M. participated in the interpretation of data and in finalising the paper. All authors read and approved the final manuscript.

None of the authors has any conflicts of interest to declare.

References

1. Forouzanfar MH, Alexander L, Anderson HR, *et al.* (2015) Global, regional, and national comparative risk assessment of 79 behavioural, environmental and occupational, and metabolic risks or clusters of risks in 188 countries, 1990–2013: a systematic analysis for the Global Burden of Disease Study 2013. *Lancet* **386**, 2287–2323.
2. Nordic Council of Ministers (2014) *Nordic Nutrition Recommendations 2012*, 5th ed. Copenhagen: Nordic Council of Ministers.
3. US Department of Health and Human Services & US Department of Agriculture (2015) *2015–2020 Dietary Guidelines for Americans*, 8th ed. Washington, DC: US Department of Health and Human Services and US Department of Agriculture.

4. Jacobs DR Jr & Tapsell LC (2015) What an anticardiovascular diet should be in 2015. *Curr Opin Lipidol* **26**, 270–275.
5. Koch M, Nothlings U & Lieb W (2015) *A priori*-defined dietary patterns and mortality: recent insights. *Curr Opin Lipidol* **26**, 346–347.
6. Wirfalt E, Drake I & Wallstrom P (2013) What do review papers conclude about food and dietary patterns? *Food Nutr Res* **57**, 20523.
7. Moeller SM, Reedy J, Millen AE, *et al.* (2007) Dietary patterns: challenges and opportunities in dietary patterns research an Experimental Biology workshop, April 1, 2006. *J Am Diet Assoc* **107**, 1233–1239.
8. Waijers PMCM, Feskens EJM & Ocké MC (2007) A critical review of predefined diet quality scores. *Br J Nutr* **97**, 219–231.
9. Anderson JJ & Nieman DC (2016) Diet quality – the Greeks had it right! *Nutrients* **8**, 636.
10. Martinez-Gonzalez MA, Salas-Salvado J, Estruch R, *et al.* (2015) Benefits of the Mediterranean diet: insights from the PREDIMED Study. *Prog Cardiovasc Dis* **58**, 50–60.
11. de Lorgeril M, Salen P, Martin JL, *et al.* (1999) Mediterranean diet, traditional risk factors, and the rate of cardiovascular complications after myocardial infarction: final report of the Lyon Diet Heart Study. *Circulation* **99**, 779–785.
12. Estruch R, Ros E, Salas-Salvado J, *et al.* (2013) Primary prevention of cardiovascular disease with a Mediterranean diet. *N Engl J Med* **368**, 1279–1290.
13. Liyanage T, Ninomiya T, Wang A, *et al.* (2016) Effects of the Mediterranean Diet on cardiovascular outcomes – a systematic review and meta-analysis. *PLOS ONE* **11**, e0159252.
14. Martinez-Gonzalez MA & Martin-Calvo N (2016) Mediterranean diet and life expectancy; beyond olive oil, fruits, and vegetables. *Curr Opin Clin Nutr Metab Care* **19**, 401–407.
15. Adamsson V, Cederholm T, Vessby B, *et al.* (2014) Influence of a healthy Nordic diet on serum fatty acid composition and associations with blood lipoproteins – results from the NORDIET study. *Food Nutr Res* **58**, 24114.
16. Uusitupa M, Hermansen K, Savolainen MJ, *et al.* (2013) Effects of an isocaloric healthy Nordic diet on insulin sensitivity, lipid profile and inflammation markers in metabolic syndrome – a randomized study (SYSDIET). *J Intern Med* **274**, 52–66.
17. Olsen A, Egeberg R, Halkjaer J, *et al.* (2011) Healthy aspects of the Nordic diet are related to lower total mortality. *J Nutr* **141**, 639–644.
18. Roswall N, Eriksson U, Sandin S, *et al.* (2015) Adherence to the healthy Nordic food index, dietary composition, and lifestyle among Swedish women. *Food Nutr Res* **59**, 26336.
19. Kanerva N, Kaartinen NE, Schwab U, *et al.* (2014) The Baltic Sea Diet Score: a tool for assessing healthy eating in Nordic countries. *Public Health Nutr* **17**, 1697–1705.
20. Hillesund ER, Bere E, Haugen M, *et al.* (2014) Development of a New Nordic Diet score and its association with gestational weight gain and fetal growth – a study performed in the Norwegian Mother and Child Cohort Study (MoBa). *Public Health Nutr* **17**, 1909–1918.
21. Bjornara HB, Overby NC, Stea TH, *et al.* (2016) The association between adherence to the New Nordic Diet and diet quality. *Food Nutr Res* **60**, 31017.
22. Roswall N, Sandin S, Löf M, *et al.* (2015) Adherence to the healthy Nordic food index and total and cause-specific mortality among Swedish women. *Eur J Epidemiol* **30**, 509–517.
23. Roswall N, Sandin S, Scragg R, *et al.* (2015) No association between adherence to the healthy Nordic food index and cardiovascular disease amongst Swedish women: a cohort study. *J Intern Med* **278**, 531–541.
24. Kanerva N, Kaartinen NE, Rissanen H, *et al.* (2014) Associations of the Baltic Sea diet with cardiometabolic risk factors – a meta-analysis of three Finnish studies. *Br J Nutr* **112**, 616–626.
25. Larsson SC, Bergkvist L & Wolk A (2009) Long-term dietary calcium intake and breast cancer risk in a prospective cohort of women. *Am J Clin Nutr* **89**, 277–282.
26. Tektonidis TG, Åkesson A, Gigante B, *et al.* (2015) A Mediterranean diet and risk of myocardial infarction, heart failure and stroke: a population-based cohort study. *Atherosclerosis* **243**, 93–98.
27. Willett WC, Howe GR & Kushi LH (1997) Adjustment for total energy intake in epidemiologic studies. *Am J Clin Nutr* **65**, 1220S–1228S; discussion 1229S–1231S.
28. Orsini N, Bellocco R, Bottai M, *et al.* (2008) Validity of self-reported total physical activity questionnaire among older women. *Eur J Epidemiol* **23**, 661–667.
29. Charlson ME, Pompei P, Ales KL, *et al.* (1987) A new method of classifying prognostic comorbidity in longitudinal studies: development and validation. *J Chronic Dis* **40**, 373–383.
30. Quan H, Sundararajan V, Halfon P, *et al.* (2005) Coding algorithms for defining comorbidities in ICD-9-CM and ICD-10 administrative data. *Med Care* **43**, 1130–1139.
31. Trichopoulou A, Costacou T, Bamia C, *et al.* (2003) Adherence to a Mediterranean diet and survival in a Greek population. *N Engl J Med* **348**, 2599–2608.
32. Bellavia A, Tektonidis TG, Orsini N, *et al.* (2016) Quantifying the benefits of Mediterranean diet in terms of survival. *Eur J Epidemiol* **31**, 527–530.
33. VanderWeele TJ, Hernan MA & Robins JM (2008) Causal directed acyclic graphs and the direction of unmeasured confounding bias. *Epidemiology* **19**, 720–728.
34. Michaelsson K, Wolk A, Langenskiöld S, *et al.* (2014) Milk intake and risk of mortality and fractures in women and men: cohort studies. *BMJ* **349**, g6015.
35. Riserus U (2015) Healthy Nordic diet and cardiovascular disease. *J Intern Med* **278**, 542–544.
36. Adamsson V, Reumark A, Cederholm T, *et al.* (2012) What is a healthy Nordic diet? Foods and nutrients in the NORDIET study. *Food Nutr Res* **56**, 18189.
37. Amcoff E, Edberg A, Enghardt Barbieri H, *et al.* (2012) *Riksmaten – vuxna 2010–11. Livsmedels- och näringsintag bland vuxna i Sverige (Riksmaten Adults 2010–11. Food and Nutrient Intake Among Adults in Sweden)*. Uppsala: Livsmedelsverket.
38. Pedersen AN, Christensen T, Matthiessen J, *et al.* (2015) *Danskernes kostvaner 2011–2013. Hovedresultater (Dietary Habits in Denmark 2011–2013. Main Results)*. Søborg: DTU Fødevareinstituttet, Afdeling for Ernæring.
39. McKibben RA, Al Rifai M, Mathews LM, *et al.* (2016) Primary prevention of atherosclerotic cardiovascular disease in women. *Curr Cardiovasc Risk Rep* **10**, 1.
40. Gunge VB, Andersen I, Kyro C, *et al.* (2017) Adherence to a healthy Nordic food index and risk of myocardial infarction in middle-aged Danes: the diet, cancer and health cohort study. *Eur J Clin Nutr* **71**, 652–658.
41. Martinez-Gonzalez MA (2016) Benefits of the Mediterranean diet beyond the Mediterranean Sea and beyond food patterns. *BMC Med* **14**, 157.
42. Martinez-Gonzalez MA, Garcia-Lopez M, Bes-Rastrollo M, *et al.* (2011) Mediterranean diet and the incidence of cardiovascular disease: a Spanish cohort. *Nutr Metab Cardiovasc Dis* **21**, 237–244.
43. Tektonidis TG, Åkesson A, Gigante B, *et al.* (2016) Adherence to a Mediterranean diet is associated with reduced risk of heart failure in men. *Eur J Heart Fail* **18**, 253–259.
44. Tong TY, Wareham NJ, Khaw KT, *et al.* (2016) Prospective association of the Mediterranean diet with cardiovascular disease incidence and mortality and its population impact in a

- non-Mediterranean population: the EPIC-Norfolk study. *BMC Med* **14**, 135.
45. Grosso G, Mistretta A, Frigiola A, *et al.* (2014) Mediterranean diet and cardiovascular risk factors: a systematic review. *Crit Rev Food Sci Nutr* **54**, 593–610.
 46. Schwingshackl L & Hoffmann G (2014) Mediterranean dietary pattern, inflammation and endothelial function: a systematic review and meta-analysis of intervention trials. *Nutr Metab Cardiovasc Dis* **24**, 929–939.
 47. Bere E & Brug J (2010) Is the term 'Mediterranean diet' a misnomer? *Public Health Nutr* **13**, 2127–2129.
 48. Rosato V, Temple NJ, La Vecchia C, *et al.* (2017) Mediterranean diet and cardiovascular disease: a systematic review and meta-analysis of observational studies. *Eur J Nutr* (publication ahead of print version 25 November 2017).
 49. Rhee JJ, Sampson L, Cho E, *et al.* (2015) Comparison of methods to account for implausible reporting of energy intake in epidemiologic studies. *Am J Epidemiol* **181**, 225–233.
 50. Paeratakul S, Popkin BM, Kohlmeier L, *et al.* (1998) Measurement error in dietary data: implications for the epidemiologic study of the diet–disease relationship. *Eur J Clin Nutr* **52**, 722–727.